

Efficient Price Sensitivity Estimation of Path-Dependent Derivatives by Weak Derivatives

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- 2 WDM: General Discussion
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Motivation

- Development of more and more complicated financial products
 - more complex pricing
 - growing emphasis on risk management issues
 - global computation of risk figures such as VaR and CVaR
- Development of **efficient** methods for the computation of price sensitivities w.r.t. model parameters (“Greeks”)
 - Restriction: **computation time**, since, in many cases, these risk figures are not available in **closed formulas**
 - requirement of numerical methods!

Important Fact

Monte Carlo method is often the only applicable method!

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Description of the Estimation Problem

- **Derivative Price:** Let $J := \xi \cdot \mathbb{E}[L(Y)]$, where ξ denotes a deterministic discount factor, L a payoff and ϑ the parameter of interest. Then

$$\frac{d}{d\vartheta} J(\vartheta) = \xi(\vartheta) \cdot \frac{d}{d\vartheta} \mathbb{E}[L] + \mathbb{E}[L] \cdot \frac{d}{d\vartheta} \xi(\vartheta).$$

- **Complicated case:** Financial derivatives with discontinuous payoff L , e.g., $L := \mathbb{1}\{Y > K\}$
- **Mathematical problem:** Find a (random) vector-valued function g_ϑ such that

$$\nabla_\vartheta \mathbb{E}[L] = \mathbb{E}[g_\vartheta] \quad (\nabla_\vartheta := (\frac{\partial}{\partial \vartheta_1}, \dots, \frac{\partial}{\partial \vartheta_n})),$$

where the function g_ϑ is called the stochastic gradient estimator of $\nabla_\vartheta \mathbb{E}[L]$.

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Standard Market Methods

● Finite Difference (FD)

- $\frac{L(Y(\vartheta+\Delta\vartheta))-L(Y(\vartheta))}{\Delta\vartheta}$ (Forward FD Estimator)
- Market standard (because very simple)
- But **biased estimator** & **large variance** for discontinuous payoff L !

● Score Function (SF)

- Introduced by Broadie & Glasserman (1996) to overcome disappointing performance of FD estimators
- Estimator:

$$L(Y) \frac{d}{d\vartheta} \log f(Y; \vartheta),$$

where f denotes the density of Y .

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Weak Derivative Representation

- The WDM assumes the following representation:

$$\mathbb{E}_\vartheta[L(X)] = \int L(x) \mu_\vartheta(dx) \Rightarrow \frac{d}{d\vartheta} \mathbb{E}_\vartheta[L(X)] = \int L(x) \mu'_\vartheta(dx).$$

- **Main Idea:** Replace $\mu_\vartheta^{(k)}$ by one of the representations of its weak derivative.
- **One possibility: Hahn-Jordan decomposition**

$$\int L(x) \mu_\vartheta^{(k)}(dx) = c_\vartheta^{(k)} \left(\int L(x) \mu_\vartheta^{(k,+)}(dx) - \int L(x) \mu_\vartheta^{(k,-)}(dx) \right).$$

- Corresponding WD estimator:

$$\mathbf{g}_\vartheta^{(k)}(\mathbf{X}^{(k,+)}, \mathbf{X}^{(k,-)}) = \mathbf{c}_\vartheta^{(k)} (\mathbf{L}(\mathbf{X}^{(k,+)}) - \mathbf{L}(\mathbf{X}^{(k,-)})), \text{ where}$$

$\mathbf{X}^{(k,+)} \sim \mu_\vartheta^{(k,+)}$ and $\mathbf{X}^{(k,-)} \sim \mu_\vartheta^{(k,-)}$ are independent r.v.

Weak Derivative Representation

- The WDM assumes the following representation:

$$\mathbb{E}_\vartheta[L(X)] = \int L(x) \mu_\vartheta(dx) \Rightarrow \frac{d^k}{d\vartheta^k} \mathbb{E}_\vartheta[L(X)] = \int L(x) \mu_\vartheta^{(k)}(dx).$$

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Weak Derivative Representation (cont'd)

- A stronger condition for weak differentiability is **absolute continuity**.
- This condition guarantees the existence of a density f_{ϑ} .
- The weak derivative of k^{th} -order has the representation

$$\frac{\partial^k f_{\vartheta}}{\partial \vartheta^k} = c_{\vartheta}^{(k)} (f_{\vartheta}^{(k,1)} - f_{\vartheta}^{(k,2)}),$$

where $f_{\vartheta}^{(k,1)}$ and $f_{\vartheta}^{(k,2)}$ are probability densities.

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Model Setup

- System composed of a collection of independent normal r.v. $\{X_i; i = 1, \dots, n\}$ with joint p.d.f.

$$\phi_{\vartheta}(x) := \prod_{i=1}^n \phi_{i,\vartheta}(x_i), \quad \text{where} \quad \phi_{i,\vartheta}(x_i) := \frac{e^{-\frac{1}{2} \left(\frac{x_i - \mu_i(\vartheta)}{\nu_i(\vartheta)} \right)^2}}{\sqrt{2\pi} \nu_i(\vartheta)}.$$

- Notice that this collection might describe a discrete Markov process $\{\tilde{X}_i\}_{i=1}^n$ with deterministic initial value $\tilde{X}_0 = \tilde{x}_0$ and transition p.d.f.

$$\phi_{i,\vartheta}(\tilde{x}_i; \tilde{x}_{i-1}) := \frac{e^{-\frac{1}{2} \left(\frac{\tilde{x}_i - \tilde{x}_{i-1} - \tilde{\mu}_i(\vartheta)}{\tilde{\nu}_i(\vartheta)} \right)^2}}{\sqrt{2\pi} \tilde{\nu}_i(\vartheta)} \quad \text{given} \quad \tilde{X}_{i-1} = \tilde{x}_{i-1}.$$

Model Setup (cont'd)

- Notation:

$$X_{Y_i} := (X_1, \dots, X_{i-1}, Y_i, X_{i+1}, \dots, X_n),$$

$$X_{Y_{ij}} := (X_1, \dots, X_{i-1}, Y_i, X_{i+1}, \dots, X_{j-1}, Y_j, X_{j+1}, \dots, X_n) \quad (i \neq j),$$

where Y_i and Y_j are independent r.v. with p.d.f. $f_{i,\vartheta}$ and $f_{j,\vartheta}$, respectively.

- Both r.v. are independent of all X_l ($l \neq i, j$).
- The joint p.d.f. of the set of independent r.v.

$\{X_1, \dots, Y_i, \dots, X_n\}$ and $\{X_1, \dots, Y_i, \dots, Y_j, \dots, X_n\}$ is given by

$$\prod_{l=1}^{i-1} \phi_{l,\vartheta}(x_l) f_{i,\vartheta}(y_i) \prod_{l=i+1}^n \phi_{l,\vartheta}(x_l) \quad \text{and} \quad \prod_{\substack{l=1 \\ l \neq i,j}}^n \phi_{l,\vartheta}(x_l) f_{i,\vartheta}(y_i) f_{j,\vartheta}(y_j),$$

respectively.

Used Random Variables

- $X = (X_1, \dots, X_n)$, $X_i \sim N(\mu_i, \nu_i)$
- $W^\pm = (W_1^\pm, \dots, W_n^\pm)$, $W_i^\pm \sim \text{WB}(2, \pm\nu_i \sqrt{2}, \mu_i)$
(WB = Weibull Distribution)
- $M = (M_1, \dots, M_n)$, $M_i \sim \text{DM}(\mu_i, \nu_i)$
(DM = Double-Maxwell Distribution)
- $G^\pm = (G_1^\pm, \dots, G_n^\pm)$, $G_i^\pm \sim G(\mu_i, \pm\nu_i)$, with density
$$f(x; \mu, \nu) := \begin{cases} \frac{1}{2\nu} \left(\frac{x-\mu}{\nu}\right)^3 \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\nu}\right)^2\right), & x \geq \mu \\ 0, & x < \mu \end{cases}$$
- $B = (B_1, \dots, B_n)$, $B_i \sim B(\mu_i, \nu_i)$ with density
$$f(x; \mu, \nu) := \frac{1}{3\nu\sqrt{2\pi}} \left(\frac{x-\mu}{\nu}\right)^4 \exp\left(-\frac{1}{2}\left(\frac{x-\mu}{\nu}\right)^2\right), \quad \forall x \in \mathbb{R}$$

General WD estimator

Applying the WDM to $\frac{d}{d\vartheta} \xi(\vartheta) \mathbb{E}_{\vartheta}[L(X)]$ leads to the following unbiased sensitivity estimator $g_{\vartheta}^{(1)}$:

WD estimator for the 1st derivative

$$g_{\vartheta}^{(1)}(x^{(1)}) = L(X) \frac{d}{d\vartheta} \xi(\vartheta) + \xi(\vartheta) \sum_{i=1}^n \left(a_{i,\vartheta}^{(1)} \Delta L_i^0 + b_{i,\vartheta}^{(1)} \Delta L_i^1 \right)$$

$$x^{(1)} := (W^{\pm}, M, X),$$

$$\Delta L_i^0 := L(X_{W_i^+}) - L(X_{W_i^-})$$

$$\Delta L_i^1 := L(X_{M_i}) - L(X)$$

$$a_{i,\vartheta}^{(1)} := \frac{1}{\nu_i \sqrt{2\pi}} \frac{d\mu_i}{d\vartheta},$$

$$b_{i,\vartheta}^{(1)} := \frac{1}{\nu_i} \frac{d\nu_i}{d\vartheta}$$

General WD estimator (cont'd)

Applying the WDM to $\frac{d^2}{d\vartheta^2} \xi(\vartheta) \mathbb{E}_{\vartheta}[L(X)]$ leads to the following unbiased sensitivity estimator $g_{\vartheta}^{(2)}$:

WD estimator for the 2nd derivative

$$\begin{aligned}
 g_{\vartheta}^{(2)}(x^{(2)}) &= L(X) \frac{d^2}{d\vartheta^2} \xi(\vartheta) + \xi(\vartheta) \sum_{i=1}^n \left\{ a_{i,\vartheta}^{(2)} \Delta L_i^0 + b_{i,\vartheta}^{(2)} \Delta L_i^1 \right. \\
 &\quad + c_{i,\vartheta}^{(2)} (2 \Delta L_i^2 - 3 \Delta L_i^0) + d_{i,\vartheta}^{(2)} \Delta L_i^1 + e_{i,\vartheta}^{(2)} \Delta L_i^3 \\
 &\quad \left. + \sum_{j=1, j \neq i}^n a_{i,\vartheta}^{(k-1)} a_{j,\vartheta}^{(k-1)} \Delta L_{ij}^0 + b_{i,\vartheta}^{(k-1)} b_{j,\vartheta}^{(k-1)} \Delta L_{ij}^1 \right\} \\
 &\quad + 2 \frac{d^{k-1}}{d\vartheta^{k-1}} \xi(\vartheta) \sum_{i=1}^n a_{i,\vartheta}^{(k-1)} \Delta L_i^0 + b_{i,\vartheta}^{(k-1)} \Delta L_i^1
 \end{aligned}$$

General WD estimator (cont'd)

Applying the WDM to $\frac{d^2}{d\vartheta^2} \xi(\vartheta) \mathbb{E}_\vartheta[L(X)]$ leads to the following unbiased sensitivity estimator $g_\vartheta^{(2)}$:

WD estimator for the 2nd derivative

$$g_\vartheta^{(2)}(x^{(2)}) = L(X) \frac{d^2}{d\vartheta^2} \xi(\vartheta) + \xi(\vartheta) \sum_{i=1}^n \left\{ a_{i,\vartheta}^{(2)} \Delta L_i^0 + b_{i,\vartheta}^{(2)} \Delta L_i^1 + c_{i,\vartheta}^{(2)} (2 \Delta L_i^2 - 3 \Delta L_i^0) + d_{i,\vartheta}^{(2)} \Delta L_i^1 + e_{i,\vartheta}^{(2)} \Delta L_i^3 \right\}$$

General WD estimator (cont'd)

WD estimator for the 2nd derivative (cont'd)

$$x^{(2)} := (W^\pm, M, X, G^\pm, B),$$

$$\Delta L_i^2 := L(X_{G_i^+}) - L(X_{G_i^-}), \quad \Delta L_i^3 := 2L(X) + 3L(X_{B_i}) - 5L(X_{M_i}),$$

$$\Delta L_{ij}^0 := L(X_{W_{ij}^+}) - L(X_{W_{ij}^-}), \quad \Delta L_{ij}^1 := L(X_{M_{ij}}) - L(X),$$

$$a_{i,\vartheta}^{(2)} := \frac{1}{\nu_i \sqrt{2\pi}} \frac{d^2 \mu_i}{d\vartheta^2},$$

$$b_{i,\vartheta}^{(2)} := \left(\frac{1}{\nu_i} \frac{d\mu_i}{d\vartheta} \right)^2$$

$$c_{i,\vartheta}^{(2)} := \sqrt{\frac{2}{\pi}} \frac{1}{\nu_i^2} \frac{d\mu_i}{d\vartheta} \frac{d\nu_i}{d\vartheta},$$

$$d_{i,\vartheta}^{(2)} := \frac{1}{\nu_i} \frac{d^2 \nu_i}{d\vartheta^2},$$

$$e_{i,\vartheta}^{(2)} := \left(\frac{1}{\nu_i} \frac{d\nu_i}{d\vartheta} \right)^2.$$

General WD estimator (cont'd)

- $n_a := \max\{i \in \{1, \dots, n\} : a_{i,\vartheta}^{(1)} \neq 0\}$, $n_b := \max\{i \in \{1, \dots, n\} : b_{i,\vartheta}^{(1)} \neq 0\}$
- We make the following observations:

$$a_{i,\vartheta}^{(1)} = 0 \Rightarrow a_{i,\vartheta}^{(2)}, b_{i,\vartheta}^{(2)}, c_{i,\vartheta}^{(2)} = 0 \quad \text{and} \quad b_{i,\vartheta}^{(1)} = 0 \Rightarrow c_{i,\vartheta}^{(2)}, d_{i,\vartheta}^{(2)}, e_{i,\vartheta}^{(2)} = 0.$$

Reformulated WD estimators

$$g_{\vartheta}^{(1)}(x^{(1)}) = L(X) \frac{d}{d\vartheta} \xi(\vartheta) + \xi(\vartheta) \left(\sum_{i=1}^{n_a} a_{i,\vartheta}^{(1)} \Delta L_i^0 + \sum_{i=1}^{n_b} b_{i,\vartheta}^{(1)} \Delta L_i^1 \right)$$

$$g_{\vartheta}^{(2)}(x^{(2)}) = L(X) \frac{d^2}{d\vartheta^2} \xi(\vartheta) + \xi(\vartheta) \left[\sum_{i=1}^{n_a} (a_{i,\vartheta}^{(2)} \Delta L_i^0 + b_{i,\vartheta}^{(2)} \Delta L_i^1) \right. \\ \left. + \sum_{i=1}^{n_b} (d_{i,\vartheta}^{(2)} \Delta L_i^1 + e_{i,\vartheta}^{(2)} \Delta L_i^3) + \sum_{i=1}^{n_a \wedge n_b} c_{i,\vartheta}^{(2)} (2 \Delta L_i^2 - 3 \Delta L_i^0) \right]$$

General WD estimator (cont'd)

Remark

- If $n_a, n_b \ll n$, then the extra computational cost for the WD estimator is small compared to the SF estimator.
- The magnitude of n_a and n_b depend only on the concrete model and its model parameter ϑ .
- Fortunately, important price sensitivities of models used to price equity and FX derivatives have $n_a, n_b = 1$, e.g., Delta, Gamma and Theta. The BS and CEV models, considered here, are such models that have this advantageous property.

WD estimator in the BS model

WD estimator for Greeks

$$g_{\vartheta}^{(1)}(x^{(1)}) = (-1)^{q(\vartheta)} e^{-rT} \left[\sum_{i=1}^{n_a} (a_{i,\vartheta}^{(1)} \Delta L_i^0 + \sum_{i=1}^{n_b} b_{i,\vartheta}^{(1)} \Delta L_i^1) - \frac{drT}{d\vartheta} L(X) \right],$$

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$$q(\vartheta) := \mathbf{1}\{\vartheta = T\},$$

$$a_{1,S_0}^{(1)} = \frac{1}{\nu S_0 \sqrt{2\pi}}, \quad b_{1,S_0}^{(1)} = 0, \quad a_{i \geq 2, S_0}^{(1)} = 0, \quad b_{i \geq 2, S_0}^{(1)} = 0,$$

$$a_{1,S_0}^{(2)} = -\frac{1}{\nu S_0^2 \sqrt{2\pi}}, \quad b_{1,S_0}^{(2)} = \frac{1}{\nu^2 S_0^2}, \quad a_{i \geq 2, S_0}^{(2)} = 0, \quad b_{i \geq 2, S_0}^{(2)} = 0,$$

$$a_{i,\sigma}^{(1)} = -\sqrt{\frac{\Delta t}{2\pi}}, \quad b_{i,\sigma}^{(1)} = \frac{1}{\sigma}, \quad a_{i,r}^{(1)} = \frac{1}{\sigma} \sqrt{\frac{\Delta t}{2\pi}}, \quad b_{i,r}^{(1)} = 0,$$

$$a_{1,T}^{(1)} = \frac{r - \sigma^2/2}{\nu \sqrt{2\pi}}, \quad b_{1,T}^{(1)} = \frac{1}{2\Delta t}, \quad a_{i \geq 2, T}^{(1)} = 0, \quad b_{i \geq 2, T}^{(1)} = 0.$$

WD estimator in the CEV model

WD estimator for Greeks

$$g_{\vartheta}^{(1)}(x^{(1)}) = (-1)^{q(\vartheta)} e^{-rT} \left[\sum_{i=1}^{n_a} (a_{i,\vartheta}^{(1)} \Delta L_i^0 + \sum_{i=1}^{n_b} b_{i,\vartheta}^{(1)} \Delta L_i^1) - \frac{drT}{d\vartheta} L(X) \right],$$

$$g_{\vartheta}^{(2)}(x^{(2)}) = (-1)^{q(\vartheta)} e^{-rT} \left[\sum_{i=1}^{n_a} (a_{i,\vartheta}^{(2)} \Delta L_i^0 + b_{i,\vartheta}^{(2)} \Delta L_i^1) \right. \\ \left. + \sum_{i=1}^{n_b} (d_{i,\vartheta}^{(2)} \Delta L_i^1 + e_{i,\vartheta}^{(2)} \Delta L_i^3) + \sum_{i=1}^{n_a \wedge n_b} c_{i,\vartheta}^{(2)} (2 \Delta L_i^2 - 3 \Delta L_i^0) \right]$$

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Efficiency Measure

- Standard error (stderr): Precision of the mean estimate.
- Variance reduction factor $VRF = \left(\frac{\text{stderr}_{\text{BM}}}{\text{stderr}}\right)^2$. Benchmark (BM) is SFM.
- Computational cost: Measured by the number of updates of the asset prices S_j 's, $j = i, \dots, n$ and $i = 1, \dots, n^*$, $n^* \in \{n_a, n_b\}$.
 - The following ratio indicates a deterioration of performance:

$$\frac{\# \text{Calculation } S_j \text{'s}_{\text{original}} + \# \text{WDVariables} \times \# \text{Calculation } S_j \text{'s}_{\text{additional}}}{\# \text{Calculation } S_j \text{'s}_{\text{original}}}$$

- $\# \text{WDVariables}$ denotes how many of the $Y \in \{W^\pm, M, G^\pm, B\}$ are involved in the WD estimator.
- Better efficiency measure for an estimator: *Divide the VRF by this ratio.*

$$\text{AON Call: } L(S_T) = S_T \mathbb{1}\{S_T > K\}$$

	FD	WD ^U	WD		FD	WD ^U	WD
$\Delta:K=80$	5.E-02	18	18	$\Delta:K=95$	8.E-02	23	23
$\Delta:K=90$	2.E-02	21	21	$\Delta:K=100$	3.E-02	27	26
$\Delta:K=100$	1.E-02	94	94	$\Delta:K=102$	2.E-02	51	44
$\Delta:K=110$	9.E-03	42	42	$\Delta:K=106$	2.E-02	247	125
$\Delta:K=120$	9.E-03	13	13	$\Delta:K=110$	2.E-02	18	17
$\Delta:K=150$	1.E-02	7	7	$\Delta:K=120$	3.E-02	8	8
$\Gamma:K=80$	2.E-04	14	4	$\Gamma:K=95$	2.E-03	12	6
$\Gamma:K=90$	7.E-05	11	3	$\Gamma:K=100$	1.E-03	12	4
$\Gamma:K=100$	4.E-05	27	2	$\Gamma:K=102$	8.E-04	16	3
$\Gamma:K=110$	3.E-05	14	2	$\Gamma:K=106$	6.E-04	33	3
$\Gamma:K=120$	3.E-05	7	2	$\Gamma:K=110$	6.E-04	9	3
$\Gamma:K=150$	8.E-05	6	3	$\Gamma:K=120$	2.E-03	9	6

Table: VRF: BS Model

Table: VRF: CEV Model

$$\text{AON Call: } L(S_T) = S_T \mathbb{1}\{S_T > K\}$$

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$\Delta:K=80$	3.E-02	6	6	$\Delta:K=95$	4.E-02	6	6
$\Delta:K=90$	1.E-02	7	7	$\Delta:K=100$	2.E-02	7	7
$\Delta:K=100$	5.E-03	31	31	$\Delta:K=102$	1.E-02	13	11
$\Delta:K=110$	5.E-03	14	14	$\Delta:K=106$	1.E-02	62	31
$\Delta:K=120$	5.E-03	4	4	$\Delta:K=110$	1.E-02	5	4
$\Delta:K=150$	5.E-03	2	2	$\Delta:K=120$	2.E-02	2	2
$\Gamma:K=80$	1.E-04	4	1	$\Gamma:K=95$	1.E-03	2	1
$\Gamma:K=90$	4.E-05	3	1	$\Gamma:K=100$	5.E-04	2	1
$\Gamma:K=100$	2.E-05	7	0.5	$\Gamma:K=102$	4.E-04	2	0.5
$\Gamma:K=110$	2.E-05	4	0.5	$\Gamma:K=106$	3.E-04	5	0.5
$\Gamma:K=120$	2.E-05	2	0.5	$\Gamma:K=110$	3.E-04	1	0.5
$\Gamma:K=150$	4.E-05	2	1	$\Gamma:K=120$	1.E-03	1	1

Table: VRF/Ratio: BS Model

Table: VRF/Ratio: CEV Model

Single Barrier AON call: $L(S) = S_T \mathbb{1}\{\min_{i=1,\dots,250} S_i > K\}$

	FD	WD ^U	WD
$\Gamma:K=80$	3	125	55
$\Gamma:K=85$	2	65	29
$\Gamma:K=90$	0.8	32	15
$\Gamma:K=95$	0.3	13	6
$\Gamma:K=100$	0.1	7	2
$\Gamma:K=102$	0.2	5	3
$\kappa:K=80$	0.6	183	85
$\kappa:K=85$	0.4	112	51
$\kappa:K=90$	0.3	77	35
$\kappa:K=95$	0.2	56	25
$\kappa:K=100$	0.8	43	19
$\kappa:K=102$	0.1	36	17

Table: VRF: BS Model

	FD	WD ^U	WD
$\Gamma:K=90$	102	261	118
$\Gamma:K=95$	19	49	22
$\Gamma:K=97$	9	23	11
$\Gamma:K=98$	6	14	7
$\Gamma:K=100$	2	7	2
$\Gamma:K=101$	4	6	4
$\kappa:K=90$	0.9	347	167
$\kappa:K=95$	0.3	106	50
$\kappa:K=97$	0.3	75	35
$\kappa:K=98$	0.2	61	30
$\kappa:K=100$	0.5	48	22
$\kappa:K=101$	0.05	35	19

Table: VRF: CEV Model

Single Barrier AON call: $L(S) = S_T \mathbb{1}\{\min_{i=1,\dots,250} S_i > K\}$

	FD	WD ^U	WD
$\Gamma:K=80$	2	31	14
$\Gamma:K=85$	1	16	7
$\Gamma:K=90$	0.4	8	4
$\Gamma:K=95$	0.2	3	2
$\Gamma:K=100$	0.05	2	0.5
$\Gamma:K=102$	0.05	1	1
$\kappa:K=80$	0.3	0.5	0.2
$\kappa:K=85$	0.2	0.3	0.1
$\kappa:K=90$	0.2	0.2	0.1
$\kappa:K=95$	0.1	0.1	0.1
$\kappa:K=100$	0.4	0.1	0.05
$\kappa:K=102$	0.05	0.1	0.05

Table: VRF/Ratio: BS Model

	FD	WD ^U	WD
$\Gamma:K=90$	51	52	24
$\Gamma:K=95$	9	10	4
$\Gamma:K=97$	5	5	2
$\Gamma:K=98$	3	3	1
$\Gamma:K=100$	1	1	0.4
$\Gamma:K=101$	2	1	1
$\kappa:K=90$	0.5	3	1
$\kappa:K=95$	0.2	0.8	0.4
$\kappa:K=97$	0.2	0.6	0.3
$\kappa:K=98$	0.1	0.5	0.2
$\kappa:K=100$	0.3	0.4	0.2
$\kappa:K=101$	0.03	0.3	0.2

Table: VRF/Ratio: CEV Model

Fixed Lookback call:

$$L(S) = (\max_{i=1,\dots,250} S_i - K) \mathbb{1}\{\max_{i=1,\dots,250} S_i > K\}$$

	FD	WD ^U	WD
$\Gamma:K=110$	97	605	98
$\Gamma:K=120$	52	450	73
$\Gamma:K=130$	31	341	56
$\Gamma:K=150$	15	225	38
$\kappa:K=110$	473	824	137
$\kappa:K=120$	211	566	96
$\kappa:K=130$	102	425	71
$\kappa:K=150$	34	275	50

Table: VRF: BS Model

	FD	WD ^U	WD
$\Gamma:K=100$	1778	1340	212
$\Gamma:K=105$	368	793	130
$\Gamma:K=110$	141	505	83
$\Gamma:K=115$	67	334	57
$\kappa:K=100$	1812	1918	307
$\kappa:K=105$	787	996	169
$\kappa:K=110$	255	592	101
$\kappa:K=115$	88	395	67

Table: VRF: CEV Model

Fixed Lookback call:

$$L(S) = (\max_{i=1, \dots, 250} S_i - K) \mathbb{1}\{\max_{i=1, \dots, 250} S_i > K\}$$

	FD	WD ^U	WD
$\Gamma:K=110$	49	151	25
$\Gamma:K=120$	26	113	18
$\Gamma:K=130$	16	85	14
$\Gamma:K=150$	8	56	10
$\kappa:K=110$	236	7	1
$\kappa:K=120$	106	4	0.8
$\kappa:K=130$	51	3	0.6
$\kappa:K=150$	17	2	0.4

Table: VRF/Ratio: BS Model

	FD	WD ^U	WD
$\Gamma:K=100$	889	268	42
$\Gamma:K=105$	184	159	26
$\Gamma:K=110$	71	101	17
$\Gamma:K=115$	34	67	11
$\kappa:K=100$	906	15	2
$\kappa:K=105$	394	8	1
$\kappa:K=110$	128	5	0.8
$\kappa:K=115$	44	3	0.5

Table: VRF/Ratio: CEV Model

- 1 Introduction
- 2 WDM: General Discussion
- 3 WDM: Models with Gaussian Transition Laws
- 4 Numerical Results
- 5 Summary**

Summary

- Derivation of an unbiased WD sensitivity estimator in a Gaussian model framework
 - Valid for a large class of single-factor pricing models and path-dependent payoffs in use.
- From this general estimator we derived a WD estimator for all Greeks in the BS and CEV framework, respectively.
- Results of our simulation study
 - Coupled WD estimator had uniformly lower variance than the FD and SF estimator.
 - If the computational cost is taken into account, however, then only the Greeks with $n_a, n_b \ll n$, i.e., Δ , Γ and Θ , are more efficient than the standard methods. For κ and Θ this is not true any more.

Important Fact

WD estimator does not depend on the particular payoff but only on the underlying pricing model.

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Authors of Related Literature

- **Price Sensitivity:**

- Brodie, Glasserman (Pioneers in Estimation of Price Sensitivities)
- Heidergott (Pricing of American Plain-Vanilla Call by Stochastic Optimization)

- **WDM:**

- Pflug (Introduction of WDM)
- Heidergott, Vazquez (Measure-Valued Differentiation)
- Billingsley (Convergence of Probability Measures)

- **Overview:**

- Fu (Summary of WDM and other approaches)

Analysis of Computational Cost

- In order to analyse the computational effort of the evaluation of L consider a stochastically recursive sequence (SRS)

$$S_l = h(S_{l-1}, X_l) \quad l = 1, \dots, n, \quad (1)$$

representing the underlying risk factor such as the asset price, where $S_0 = s_0$ is a deterministic start value, h is a measurable state-transition mapping and X_l denotes a random input variable distributed according to the Normal distribution.

- We point out that the nominal path $S(X)$ is generated from X , i.e.,

$$(X_1, \dots, X_n) \mapsto (S_1, \dots, S_n).$$

- It is obvious that all the S_l 's are calculated and hence n calculations are needed.

Analysis of Computational Cost (cont'd)

- Note that the nominal path $S(X)$ and the perturbed path $S(X_{Y_i})$, $i = 1, \dots, n^*$, are equal up to state S_{i-1} and differ from state S_i onwards, i.e., $n - i + 1$ states will change.
- This fact can be written as follows:

$$\begin{pmatrix} Y_1 & X_2 & \dots & X_{n^*} & \dots & X_n \\ X_1 & Y_2 & \dots & X_{n^*} & \dots & X_n \\ \vdots & \vdots & \ddots & \vdots & & \vdots \\ X_1 & X_2 & \dots & Y_{n^*} & \dots & X_n \end{pmatrix} \mapsto \begin{pmatrix} S_1^{Y_1} & S_2^{Y_1} & \dots & S_{n^*}^{Y_1} & \dots & S_n^{Y_1} \\ S_1 & S_2^{Y_2} & \dots & S_{n^*}^{Y_2} & \dots & S_n^{Y_2} \\ \vdots & \vdots & \ddots & \vdots & & \vdots \\ S_1 & S_2 & \dots & S_{n^*}^{Y_{n^*}} & \dots & S_n^{Y_{n^*}} \end{pmatrix}$$

- How many of the S_j 's, $j = i, \dots, n$, need to be explicitly updated for the evaluation of the payoff L ?
- If we assume that all the S_j 's need to be updated, then the answer is

$$\frac{1}{2} n^* (2n - n^* + 1).$$

Reduction Factors

	Table BS		Table CEV	
	FD	WD ^U	FD	WD ^U
Δ	2	3	2	4
Γ	2	4	2	7

Table: Reduction factors of VRF's for European AON call

	Table BS		Table CEV	
	FD	WD ^U	FD	WD ^U
Γ	2	4	2	5
κ	2	378	2	126

Table: Reduction factors of VRF's for European single barrier AON call

Reduction Factors (cont'd)

	Table BS		Table CEV	
	FD	WD ^U	FD	WD ^U
Γ	2	4	2	5
κ	2	378	2	126

Table: Reduction factors of VRF's for European fixed Lookback call

Weakly Differentiability

Definition (Weakly Differentiability)

Let $\vartheta, \vartheta + \Delta\vartheta \in V$. If μ_ϑ is a family of elements of $\mathcal{P}(\mathbb{R}^n)$, we say that μ_ϑ is weakly differentiable if there exists a finite signed measure $\mu'_\vartheta : \mathfrak{B}(\mathbb{R}^n) \mapsto \mathbb{R}$ such that for all $L \in C_b(\mathbb{R}^n)$

$$\lim_{\Delta\vartheta \rightarrow 0} \frac{1}{\Delta\vartheta} \left(\int L(x) \mu_{\vartheta + \Delta\vartheta}(dx) - \int L(x) \mu_\vartheta(dx) \right) = \int L(x) \mu'_\vartheta(dx).$$

- Note that the above identity is for all $L \in C_b(\mathbb{R}^n)$ equivalent to

$$\frac{d}{d\vartheta} \int L(x) \mu_\vartheta(dx) = \int L(x) \mu'_\vartheta(dx).$$

Representation of Weak Derivative

Definition (Representation of Weak Derivative)

Let $\mu_{\vartheta}, \vartheta \in V$, be a family of elements of $\mathcal{P}(\mathbb{R}^n)$. We call a triple $(c_{\vartheta}^{(k)}, \mu_{\vartheta}^{(k,1)}, \mu_{\vartheta}^{(k,2)})$ consisting of a constant and two probability measures a *representation of the weak derivative* of k^{th} -order of μ_{ϑ} if μ_{ϑ} is k -times weakly differentiable at each ϑ and for all $L \in C_b(\mathbb{R}^n)$ it holds that

$$\int L(x) \mu_{\vartheta}^{(k)}(dx) = c_{\vartheta}^{(k)} \left(\int L(x) \mu_{\vartheta}^{(k,1)}(dx) - \int L(x) \mu_{\vartheta}^{(k,2)}(dx) \right).$$

Remark to Discontinuous Payoffs

Remark

Assume that the sequence $\{\mu_{\Delta\vartheta}^{(k)} := \frac{\mu_{\vartheta+\Delta\vartheta}^{(k)} - \mu_{\vartheta}^{(k)}}{\Delta\vartheta} : \Delta\vartheta \in V \setminus \{0\}\}$ converges weakly to $\mu_{\vartheta}^{(k)}$. If for $L \in \mathcal{L}^1(\mu_{\Delta\vartheta}^{(k)} : \Delta\vartheta \in V \setminus \{0\})$ we denote by D_L the set of points at which L is discontinuous, then for each bounded L , such that $\mu_{\vartheta}^{(k)}(D_L) = 0$, the identity for the representation of weak derivatives is still true. For a proof of this result when the sequence $\{\mu_{\Delta\vartheta}^{(k)}\}_{\Delta\vartheta} \subset \mathcal{P}(\mathbb{R}^n)$ see, e.g., Billingsley (1999). The extension to the case $\{\mu_{\Delta\vartheta}^{(k)}\}_{\Delta\vartheta} \subset \mathcal{M}(\mathbb{R}^n)$ is straightforward to prove.

BS Model

- The asset price S_T is described in the BS model by

$$S_T = S_t e^{(r-\sigma^2/2)(T-t)+\sigma\sqrt{T-t}Z} \quad (0 \leq t < T),$$

where $Z \sim N(0, 1)$, S_t is the current asset price, r and σ are constant.

- Divide the horizon $[0, T]$ into n equal time intervals each of length Δt : $0 = t_0 < t_1 < \dots < t_n = T$.
- S_{t_i} at times t_i with known initial asset price S_0 are generated by

$$S_{t_i} = S_{t_{i-1}} e^{\tilde{\mu} + \tilde{\nu} Z_i} \quad (i = 1, \dots, n),$$

where Z_1, \dots, Z_n is a sequence of independent, standard normal variables, $\tilde{\mu} := (r - \sigma^2/2) \Delta t$ and $\tilde{\nu} := \sigma \sqrt{\Delta t}$.

CEV Model

- Dynamics of asset price movements:

$$d S_t = r S_t dt + \sigma S_t^\gamma d W_t, \quad (2)$$

where W is a standard Brownian motion.

- The elasticity parameter γ was originally negative but was extended to include positive values.
- Euler approximation of (2)

$$S_{t_i} = S_{t_{i-1}} + \tilde{\mu}_i + \tilde{\nu}_i Z_i \quad (i = 1, \dots, n), \quad (3)$$

where $\tilde{\mu}_i := r S_{t_{i-1}} \Delta t$ and $\tilde{\nu}_i := \sigma S_{t_{i-1}}^\gamma \sqrt{\Delta t}$.

Sensitivities

- Risk-neutral pricing formula:

$$V(0, S_0) = e^{-rT} \mathbb{E}[L(S)].$$

- Assumption (A0) is satisfied with ξ given by e^{-rT} in both models.
- The first ($k = 1$) and second derivative ($k = 2$) of the price w.r.t. the parameter ϑ is as follows:

$$\begin{aligned} \frac{d^k V}{d\vartheta^k} = & -\frac{d^k rT}{d\vartheta^k} V + e^{-rT} \frac{d^k}{d\vartheta^k} \mathbb{E}[L] \\ & + \mathbf{1}\{k = 2\} \left\{ \left(\frac{d^{k-1} rT}{d\vartheta^{k-1}} \right)^2 V - 2 e^{-rT} \frac{d^{k-1} rT}{d\vartheta^{k-1}} \frac{d^{k-1}}{d\vartheta^{k-1}} \mathbb{E}[L] \right\} \end{aligned} \quad (4)$$

- In (4) $\frac{d^2 rT}{d\vartheta^2} \equiv 0$ and $\frac{d rT}{d\vartheta}$ is only non-zero for $\vartheta \in \{r, T\}$.

The Greeks

Definition (The Greeks)

DELTA: The delta (Δ) is defined as the rate of change of the option price with respect to the initial asset price, i.e.,
$$\Delta := \partial V / \partial S_0.$$

GAMMA: The gamma (Γ) is the rate of change in the delta with respect to the initial asset price, i.e., $\Gamma := \partial^2 V / \partial S_0^2$.

VEGA: The vega (κ) is the rate of change of the option price with respect to the volatility of the underlying asset, i.e., $\kappa := \partial V / \partial \sigma$.

RHO: The rho (ρ) is defined as the rate of change of the option price with respect to the interest rate, i.e., $\rho := \partial V / \partial r$.

THETA: The theta (Θ) is the negative of the rate of change of the option price with respect to the passage of time, i.e.,
$$\Theta := -\partial V / \partial T.$$